# Numerical Investigation on Asymptotic Features of Model Selection Criteria 

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## 1. INTRODUCTION

Using a synthetic example of four geostatistical models (including the true or data-generating model), this study is to investigate:

- Accuracy of BIC and KIC
- Asymptotic features of the model selection criteria, AIC, AICc, BIC, and KIC,
- Fisher information matrix's role in model selection;
- Sensitivity of posterior to prior model probability.

2. THEORETICAL BACKGROUND

The posterior distribution of the quantity of interest, $\Delta$, given a set of data $\mathbf{D}$ is:

$$
p(\Delta \mid \mathbf{D})=\sum_{k=1}^{K} p\left(\Delta \mid M_{k}, \mathbf{D}\right) p\left(M_{k} \mid \mathbf{D}\right)
$$

where $\mathbf{M}=\left(M_{1}, \ldots, M_{k}\right)$ is the set of models considered.
The posterior model probability of model $M_{k}$ is given by Bayes rule:

$$
p\left(M_{k} \mid \mathbf{D}\right)=\frac{p\left(\mathbf{D} \mid M_{k}\right) p\left(M_{k}\right)}{\sum_{l=1}^{K} p\left(\mathbf{D} \mid M_{l}\right) p\left(M_{l}\right)}
$$

where

$$
p\left(\mathbf{D} \mid M_{k}\right)=\int p\left(\mathbf{D} \mid \boldsymbol{\theta}_{k}, M_{k}\right) p\left(\boldsymbol{\theta}_{k} \mid M_{k}\right) d \boldsymbol{\theta}_{k}
$$

The likelihood function can be evaluated using either a simple Monte Carlo method via

$$
\hat{p}\left(\mathbf{D} \mid M_{k}\right)=\frac{1}{N} \sum_{i=1}^{N} p\left(\mathbf{D} \mid \boldsymbol{\theta}_{k}^{(i)}, M_{k}\right)
$$

or the Laplace approximation via model selection criteria

$$
p\left(\mathbf{D} \mid M_{k}\right)=\exp \left(-\frac{1}{2} I C_{k}\right)
$$

where IC represents $A I C, A I C C, B I C$, and $K I C$

$$
\begin{aligned}
A I C_{k} & =-2 \ln \left[L\left(\hat{\boldsymbol{\theta}}_{k} \mid \mathbf{D}\right)\right]+2 N_{k} \\
A I C C_{k} & =-2 \ln \left[L\left(\hat{\boldsymbol{\theta}}_{k} \mid \mathbf{D}\right)\right]+2 N_{k}+\frac{2 N_{k}\left(N_{k}+1\right)}{N-N_{k}-1} \\
B I C_{k} & =-2 \ln \left[L\left(\hat{\boldsymbol{\theta}}_{k} \mid \mathbf{D}\right)\right]+N_{k} \ln N \\
K I C_{k} & =-2 \ln \left[L\left(\hat{\boldsymbol{\theta}}_{k} \mid \mathbf{D}\right)\right]+N_{k} \ln N-2 \ln p\left(\hat{\boldsymbol{\theta}}_{k}\right)-N_{k} \ln 2 \pi+\ln \left|\overline{\mathbf{F}}_{k}\right| \\
\bar{F}_{k j} & =\frac{1}{N} F_{k j}=-\left.\frac{1}{N} \frac{\partial^{2} \ln \left[L\left(\hat{\boldsymbol{\theta}}_{k} \mid \mathbf{D}\right)\right]}{\partial \theta_{k i} \partial \theta_{k j}}\right|_{\boldsymbol{\theta}_{k}=\hat{\theta}_{k}} \quad \text { Ye et al., 2008) }
\end{aligned}
$$

3. THE SYNTHETIC EXAMPLE

Random data D consists of deterministic trend $\mu$ and random residual $R$ : $D=\mu+R$


True and Alternative Models
$M_{0}: \boldsymbol{\mu}(x, y)=a_{0}+a_{1} x+a_{2} y+a_{3} y^{2}$
Three alternative models, $M_{1}, M_{2}$, and $M_{3}$
$M_{1}: \boldsymbol{\mu}(x, y)=a_{0}+a_{1} x+a_{2} y+a_{3} x y+a_{4} y^{2}$
$M_{2}: \boldsymbol{\mu}(x, y)=a_{0}+a_{1} x y+a_{2} y+a_{3} y^{2}$
$M_{3}: \boldsymbol{\mu}(x, y)=a_{0}+a_{1} x+a_{2} y$
4. RESULTS AND DISCUSSIONS

### 4.1 Asymptotic Investigation

The asymptotic numerical investigation is conducted by gradually increasing the number of calibration data, $\mathbf{D}$, from 200 to 1,700. or each experiment, following Ye et al. (2004), the four models are calibrated, and the four criteria and corresponding model probabilities are calculated
4.2 Accuracy of BIC and KIC

Questions:
When does KIC approach to BIC?
How accurate is the posterior model probability calculated using BIC, instead of KIC?



Red (MC), Blue (KIC), Green (BIC) Taking $M_{3}$ as an example, KIC asymptotcally reduces to BIC , Bu KIC is more accurate than BIC for $K I C$ is more accurate than BIC for approximating the model
likelihood function and for calculating model probabilities.
4.3 Asymptotic Features of BIC and KIC



The asymptotic behavior of the difference of BIC between $M_{1}$ and $M_{3}$ is both determined by the model fitting term and the complexity penalty term. The in $\mid \overline{F_{i}}$, term of $K I C$ imposes more penalties to the complex model

### 4.4 Fisher Information Matrix

In $\left|\overline{\mathbf{F}}_{k}\right|$ is the major reason that $K I C$ is more accurate than BIC or approximation the model likelihood function.
In $\left|\overline{\mathbf{F}}_{k}\right|$ values decrease when the number of calibration data increases and its values follow the same order, $M_{1}>M_{2}>M_{0}>M_{3}$


The $\ln \left|\overline{\mathbf{F}}_{k}\right|$ term
discriminates alternative models based on not only model complexity but also model structure by KIC.
4.5 Consistency property of BIC and KIC

Over all the numerical experiments, $M_{2}$ receives negligible probability, due to its most deviation from the true model $M_{0}$ When the true model is included, KIC selects the true model in al experiments by assigning overwhelmingly large probability to it
While AICC selects $M_{1}$ for the experiment of $N=300$



Red $\left(M_{0}\right)$, Green $\left(M_{1}\right)$ and Blue $\left(M_{3}\right)$

### 4.6 Asymptotic Features of $\operatorname{AICC}$

When the true model is excluded, $A I C c$ considers $M_{1}$ as the bes model for all the experiments. The asymptotic behavior of the difference of $A I C C$ is solely determined by the model fitting term.

4.7 Sensitivity of Posterior Model Probability

The sensitivity of posterior to prior model probability is measured by the range between the maximum and minimum measured by the range between the maximum and minim smaller sensitivity).
For both AICC- and KIC-based probabilities, sensitivity of posterior to prior model probability is irrelevant to the number of calibration data.
The sensitivity is inversely related to the $\Delta K I C$ values because small $\Delta K I C$ values give large range of the posterior model probability (Ye et al., 2005)






Sensitivity of posterior probability of (a) $M_{0}$ and (b) $M_{\text {( }}$ (when the true model is included) and (c) $M_{1}$ and (d) $M_{3}$ (when the true model is excluded) for the twelve numerical experiments.

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